

Inflammatory potential of the diet and risk of gastric cancer in the European Prospective Investigation into Cancer and Nutrition (EPIC) study

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Short running head: Inflammatory potential of diet and gastric cancer.

Abbreviations used: CGC: Cardia gastric cancer CI: Confidence interval CRP: C-reactive protein DII: Dietary Inflammatory Index EPIC: European Prospective Investigation into Cancer and Nutrition FFQ: Food-frequency questionnaire

- GC: Gastric cancer
- GEJ: Gastro-esophageal junction
- 24hDR: 24-hour Dietary recall
- HR: Hazard ratio
- IARC: International Agency for Research on Cancer
- ICD-O: International Classification of Diseases for Oncology
- ISD: Inflammatory Score of the Diet
- LR: Likelihood ratio
- NCC: Non-cardia cancer
- OR: Odds ratio
- SD: Standard deviation

- 1 ABSTRACT
- 2

Background: Chronic inflammation plays a critical role in the pathogenesis of the two major types of
 gastric cancer. Several foods, nutrients, and non-nutrient food components seem to be involved in the
 regulation of chronic inflammation.

6 **Objective**: To assess the association between the inflammatory potential of the diet and the risk of

7 gastric carcinoma, overall and for the two major subsites: cardia cancers and non-cardia cancers.

8 **Design**: A total 476160 subjects (30% males, 70% females) from the European Investigation into

9 Cancer and Nutrition (EPIC) study were followed for 14 years, during which 913 incident cases of

10 gastric carcinoma were identified, including 236 located in the cardia, 341 in the distal part of the

11 stomach (non-cardia), and 336 with overlapping or unknown tumor site. The dietary inflammatory

12 potential was assessed by means of an inflammatory score of the diet (ISD), calculated using 28

13 dietary components and their corresponding inflammatory scores. The association between the ISD

14 and gastric cancer risk was estimated by hazard ratios (HR) and 95%-confidence intervals (CI)

15 calculated by multivariate Cox regression models adjusted for confounders.

16 **Results**: The inflammatory potential of diet was associated with an increased risk of gastric cancer.

17 The HR (95% CI) for each increase in one standard deviation of the ISD were 1.25 (1.12, 1.39) for all

18 gastric cancers, 1.30 (1.06, 1.59) for cardia cancers, and 1.07 (0.89, 1.28) for non-cardia cancers. The

19 corresponding values for the highest compared to the lowest quartiles of the ISD were 1.66 (1.26,

20 2.20), 1.94 (1.14, 3.30), and 1.07 (0.70, 1.70) respectively.

Conclusions: Our results suggest that low-grade chronic inflammation induced by the diet may be
 associated with gastric cancer risk. This pattern seems to be more consistent for gastric carcinomas
 located in the cardia than for those located in the distal stomach.

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25 Keywords: gastric cancer, nutrition, chronic inflammation, inflammatory score of the diet,

26 prospective studies.

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#### 28 INTRODUCTION

Gastric cancer (GC) is the fifth most common cancer and the third cause of death from cancer worldwide (1). Although often considered as a single entity, GC can be classified into two topographical subsites: cardia gastric cancers (CGC) arising at the area closest to the esophagus, and those arising in the distal parts of the stomach (non-cardia cancers, NCC). These two subsites of GC display different epidemiology features; while incidence of NCC has been declining over the past decades in almost all countries, the rates of CGC have remained stable or rose in several Western countries (2).

Chronic inflammation is known to play an important role in carcinogenesis (3) and several lines of 36 evidence suggest that inflammation plays a critical role in the pathogenesis of the two major types of 37 38 GC. The carcinomas arising in the distal stomach seem to be the consequence of a multistep process 39 starting from chronic inflammatory gastritis associated with persistent H. pylori infection, which may 40 evolve towards chronic atrophy gastritis, and subsequent changes in the gastric mucosa which appear to be precursor conditions of NCC (4). The pathogenesis of CGC is less well established, but some of 41 42 its risk factors are similar to esophageal adenocarcinoma, including obesity (5) and probably gastro-43 esophageal reflux (6), two conditions associated chronic inflammation. Further evidence of the potential role of inflammation on gastric carcinogenesis comes from its association with 44 polymorphisms in inflammation-related genes such as *IL1RN*, *IL1B*, and *TNF*- $\alpha$  (4,7). 45 Diet may play a role in the regulation of chronic inflammation; several foods and food components 46 47 have an impact on blood concentrations of inflammatory markers, including cytokines, chemokines, 48 acute-phase proteins, soluble adhesion molecules and cytokine receptors (8). Different 49 epidemiological studies have assessed the association between the inflammatory potential of diet, 50 measured by means of the dietary inflammatory index (DII), an index combining the intake of dietary 51 constituents and its association with well-known inflammatory markers (9), and gastro-intestinal 52 tumors (10-17). So far, only one hospital-based case-control study has addressed the association of 53 dietary inflammation with GC (17); the risk of GC more than doubled when comparing the highest

54 versus the lowest quartile of the DII. The sample size was relatively small (230 cases) and stratified 55 analyses according to anatomical site of the tumors were not performed.

56 In this paper we calculated an index to reflect the inflammatory potential of the diet (inflammatory 57 score of the diet, ISD) and assessed its association with the risk of GC in a large prospective cohort 58 from ten European countries. In addition we considered the potential role of dietary inflammation 59 separately for the two major anatomical subsites of gastric carcinoma (CGC and NCC), as well as for the two main histological types (intestinal and diffuse). 60

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#### 62 **METHODS**

#### 63 Study setting and population

64 The European Investigation into Cancer and Nutrition (EPIC) is a large prospective cohort study designed to investigate the relationships between diet, lifestyle, environmental factors and cancer. 65 Recruitment procedures and data collection of the EPIC study have been described elsewhere (18). In 66 summary, 521324 subjects, mostly aged 30 to 70 years, were recruited between 1992 and 2000 in 23 67 68 centers from ten European countries (France, Italy, Spain, United Kingdom, the Netherlands, Greece, 69 Germany, Sweden, Denmark, and Norway). Written informed consent was provided by all 70 participants. The ethical review boards from the International Agency for Research on Cancer (IARC) 71 and from all local centers approved the study. Prior to analysis, the following exclusions were made: participants with a prevalent cancer at baseline (25184), with missing follow-up information (4148), 72 lacking lifestyle or dietary information (6259), and those in the highest and lowest 1% of the 73 distribution for the ratio of energy intake to estimated energy requirement (9573). Therefore, our final 74 study population included 476160 participants (142241 men and 333919 women) (Supplemental 75 76 Figure 1). 77

#### 78 Follow-up and ascertainment of gastric cancer

79 Follow-up for incident cancer cases and assessment of vital status was provided through record 80 linkage with population cancer registries and national or regional mortality registries in most of the 81 participating countries. In France, Germany and Greece an active follow-up used a combination of 82 approaches, including cancer and pathology registries, health insurance records, and active follow-up 83 contacting participants or their next-of-kin. Cases were defined as malignant, primary incident GCs. 84 During the follow-up, a total of 1049 subjects were newly diagnosed with a malignant primary 85 cancer of the stomach (topographical code C16) according to the International Classification of 86 Diseases for Oncology, Third Edition (ICD-O-3). Among them 45 could not be classified with respect to their morphology, 31 were mesenchymal or other non-epithelial tumors, and 48 were lymphomas; 87 88 moreover there were 12 endocrine carcinomas; therefore the cases in the present study were 913 GCs 89 (epithelial tumors of the stomach excluding endocrine tumors), among which the vast majority (877) 90 were adenocarcinoma. The GC cases with code C16.0 were classified as CGC and those with codes 91 C16.1-C16.6 as NCC; the remaining cases had an overlapping tumor (C16.8) or could not be 92 classified according to their localization (C16.9). Furthermore, the GCs were classified according the 93 two main histologic types of the Lauren classification (intestinal and diffuse) based upon the morphology codes of the ICD-O (19,20). 94

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## 96 Diet, lifestyle and anthropometric information

A lifestyle questionnaire, anthropometric measurements using standardized procedures, and a 97 blood sample were collected at recruitment. The questionnaire included information on medical 98 99 history, socio-demographic characteristics, the highest school level reached, detailed history of 100 smoking habits, and a four-level index of physical activity. The usual diet over the previous twelve 101 months was assessed at baseline by means of country-specific validated questionnaires. In most countries, extensive quantitative food frequency questionnaires (FFQ) or semi-quantitative FFQ were 102 103 used, though some used diet-history questionnaires or a combination of diet record and FFQ. In 104 addition, highly standardized 24-hour dietary recall (24hDR) measurements were obtained from 105 representative subsamples (5%-12%) of each EPIC cohort. These data were used to correct for systematic differences between the dietary questionnaires and to minimize measurement error (21). 106 107 Food consumption data was used to calculate energy, macro and micronutrients and other dietary

108 components using country-specific food composition databases, which had been standardized across109 countries (22).

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### 111 The Inflammatory Score of the Diet (ISD)

112 We calculated the ISD to reflect the inflammatory potential of the diet taking the DII as the starting point (11). The DII comprises 45 food items (including macro and micronutrients, other dietary 113 components and foods) that have been assigned an inflammatory weight after a literature review 114 115 according to the pro- or anti-inflammatory effect of the food. The weight reflects the association of each food item with well-known inflammatory markers (IL-1β, IL-4, IL-6, IL-10, TNF-α and CRP). 116 To calculate the ISD in the present study we used the weights as reported (9) for a set of 28 food items 117 available in the EPIC databases for all centers (Supplemental Table 1). In order to calculate the 118 119 subject's ISD each individual's food item's intake was standardized using the mean and standard deviation of our study population. These z-scores were converted to percentile scores to avoid the 120 right skewness of data, and then centered on 0 by doubling each percentile score and subtracting 1. 121 The centered percentile values were then multiplied by its respective inflammatory effect score 122 123 (weight) to obtain the food item-specific ISD, which are summed to produce the overall ISD for each 124 participant.

The procedure to calculate the ISD is similar to the DII (9), but there are slight differences. First, 125 we did not use the weight for total fat to compute the ISD because the three components of dietary fat 126 127 (saturated, mono-unsaturated, and poly-unsaturated fats) are also included; therefore, using a weight (inflammatory effect score) for total fat would overestimate the inflammatory potential of the diet. 128 Second, we used a different weight for alcohol owing to its dose-dependent effect. In the original 129 130 database (9) alcohol is considered to be anti-inflammatory (it has a negative weight); however, the negative relationship with inflammatory markers has been showed only among moderate consumers 131 (less than 30-40 g/day) (23,24) and therefore, for subjects with intake >40 g/day the weight for 132 alcohol was set to 0. Finally, a major difference with respect to the DII was that to calculate the 133 134 subject's ISD, each individual's food item's intake was standardized using the mean and standard

deviation of our study population, while the DII used the mean and standard deviation of a regional worldwide database taken as a 'referent' population. However, the purpose of our study was not to compare the inflammatory potential of diet across populations, but to assess whether the inflammatory potential of the diet was associated with cancer risk. Therefore, we gave priority to internal validity and we used the mean and standard deviation from our own population to standardize the intakes of the ISD components.

By the way the ISD is calculated, positive values indicate a more pro-inflammatory diet and negative values correspond to a more anti-inflammatory diet. However, it should be noted that the weights used to calculate the score do not have units: they are only an indicator of the inflammatory potential of particular dietary component. Therefore, the value of the ISD for an individual is not an absolute measure of the inflammatory effect of the subject's diet, but a 'relative' index that allows categorizing individuals' diets on a continuum from maximally anti-inflammatory to maximally proinflammatory.

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## 149 Statistical analysis

150 Hazard ratios (HRs) and 95% confidence intervals (CI) were estimated using Cox proportional hazards models to assess the association between GC and the inflammatory potential of diet measured 151 by the ISD. Entry time was defined as age at recruitment, and exit time was age at diagnosis (cases), 152 death, or end of follow-up, whichever occurred first. Subjects with a diagnosis of stomach cancer 153 154 other than gastric carcinoma were censored at the time of diagnosis. All models were stratified by center and age at recruitment and adjusted for sex and total energy intake. Furthermore the 155 multivariate models were adjusted for the following potential confounders: education (none/primary 156 157 not completed, primary, technical/professional school, secondary school, longer education including 158 university); smoking status and intensity (never smoker; current, 1-15 cigarettes/day; current, 16-25 cigarettes/day; current, >25 cigarettes/day; former, quit ≤10 years; former, quit 11-20 years; former, 159 quit >20 years; other smokers, including occasional smokers, exclusive smokers of cigar and/or pipe, 160 161 and smokers with unknown status and/or unknown amount smoked); body mass index (BMI, kg/m<sup>2</sup>

<25, 25.00-29.99,  $\geq$ 30); alcohol consumption and intake (by quartiles) of red meat, processed meat, 162 citrus fruit, and non-citrus fresh fruit. The model for CGC did not include the intake of red or 163 164 processed meat, while the model for NCC did not include BMI. The selection of confounders was done a priori, based upon the known risk factors of gastric cancer (both CGC and NCC) available in 165 our dataset, and associated with the inflammatory potential of the diet. Some confounders are dietary 166 167 factors and can be source of components included in the ISD; therefore, the intakes of energy, alcohol, red and processed meat, and citrus and non-citrus fresh fruit were included into the multivariate model 168 169 as the residuals of a linear regression of each dietary variable on the ISD.

170 The ISD was analyzed both as a categorical classified by quartiles (with first quartile as the reference category) and as a continuous variable, divided by its standard deviation. Trend tests across 171 quartiles of the ISD were calculated by entering the categorical variable into the model as a 172 continuous term. The nonlinearity of the effect of ISD on GC risk was assessed by adding a quadratic 173 174 term to the model with the ISD as continuous variable and comparing the likelihood of the models with and without the quadratic term by means of the likelihood ratio (LR) test. A significant p-value 175 of this test would be interpreted as departure from linearity; although this is not a formal proof of 176 linearity, a non-significant p-value was interpreted as an indication of a linear effect of the ISD on GC 177 178 risk. The LR test was also used to evaluate the significance of the interaction of ISD with other variables of interest. The homogeneity of the risks of ISD for CGC and NCC, as well as for intestinal 179 and diffuse types, was assessed by means of the Wald statistic. The heterogeneity of HRs for the ISD 180 across countries was explored using a meta-analytic random-effects model. A chi-squared test based 181 182 upon the scaled Schoenfeld residuals was used to ensure that the assumptions of proportional hazards were met. A sensitivity analysis was conducted to evaluate possible reverse causality by excluding 183 subjects with two or less years of follow-up. 184

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186 *Calibration of intakes* 

187 A linear regression calibration approach was used to improve the comparability of dietary data across
188 centers and to minimize measurement error using data from the subsample of subjects 24hDR (25).

189 Sex- and country-specific calibration models were applied to obtain individual predicted values of 190 dietary exposures; the 24hDR measurements were regressed on dietary intake from the questionnaire, 191 including in the model total energy intake, age at recruitment, center, education, smoking, BMI and physical activity. Afterward, these models were used to obtain predicted values on intake for all 192 193 participants. For zero consumption values reported in the main dietary questionnaire a zero was 194 directly imputed as the corrected value, and negative values occasionally arising after regression were set to zero as well. The predicted values (calibrated intake) of each food component were used to 195 196 calculate the ISD. A bootstrap sampling procedure was used to compute the mean and standard 197 deviation of the predicted (calibrated) intake of each food component of the ISD in our population. A total of 400 repetitions were used to ensure the stability of the estimates (25). 198

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## 200 RESULTS

During an average follow-up of 14 years, a total of 913 incident cases of GC (56% males, 44% females) were identified among the 476160 subjects of the cohort. A total of 236 tumors were CGC, 341 NCC, and 336 had overlapping or unknown tumor site; regarding the histology 645 were intestinal, 222 diffuse, and for 46 could not be classified according the Lauren classification. Overall the inflammatory potential of the diet in the whole cohort, as measured by the ISD, had a mean value of 0.38 with a standard deviation of 1.70. The range of the ISD was from -6.44 to 5.67; the median and 25<sup>th</sup> and 75<sup>th</sup> percentiles were 0.53, -0.75, and 1.65 respectively.

208 The distribution of subjects and the ISD according to the main characteristics of the population are presented in the Table 1. The women had a more pro-inflammatory diet than men, and the 209 inflammatory potential of the diet increased with age. The ISD increased with BMI, current smokers 210 had a remarkably higher ISD than never or former smokers, while no a clear pattern was shown for 211 212 the ISD with respect of education. The ISD was also positively associated with the intake of red and processed meat, and inversely associated with the intake of citrus and other fresh fruit and with 213 alcohol consumption. Although the absolute differences were often small (in the ISD scale), all of 214 215 them were statistically significant owing to the large sample size.

216 A more complete picture of the relationship between the ISD and the usual diet of the EPIC 217 population is shown in **Supplemental Table 2**. As expected from the weight (inflammatory effect 218 score) of the dietary components of the ISD, there was a strong inverse correlation between the index 219 and the intake of legumes, vegetables, fruits (all kinds), condiments and sauces, fruit juices, coffee 220 and tea, and to a lesser extent, cereal products and alcoholic beverages; this means that all these food 221 groups tend to confer anti-inflammatory capacity to the diet. On the contrary, a strong positive correlation was evident for meat and meat products (including red and processed meat), foods based 222 223 on fats and oils, and sugar and confectionery; according to this, diets rich in these foods tend to have a 224 higher inflammatory potential.

225 The association of the inflammatory potential of the diet with GC, overall and according to the location of the tumor and the histological type, is presented in the **Table 2**. There was an increasing 226 risk of GC with higher values of the ISD, evident both for the categorized and the continuous variable. 227 228 Part of the effect of the ISD can be explained by the other risk factors of GC, but an independent association with the ISD remains after adjusting for the relevant confounders. For GC, a significant 229 230 HR was observed for each quartile of the ISD as compared with the lowest, with a significant trend. For the highest quartile there was a 66% increased risk of GC (HR 1.66, 95% CI 1.26, 2.20), and the 231 232 risk of GC significantly increased by 25% (HR 1.25, CI 1.12, 1.39) for each SD increase in the score. This association with the ISD seemed to be more consistent for tumors located in the cardia than 233 234 for those located in the distal stomach. The adjusted HRs for one SD increase in the ISD were 1.30 (1.06, 1.59) and 1.07 (0.89, 1.28) for the CGC and the NCC respectively. Despite the apparently 235 236 different effect of ISD by anatomical site, no heterogeneity of the association was observed, with nonsignificant Wald test comparing the HRs of CGC versus NCC (p-value 0.08). It is worth noting that, 237 238 contrary to CGC, the significant HR for NCC in the basic model became no-significant in the 239 multivariate model. Stepwise analysis (results not shown) showed that only the inclusion of smoking 240 and/or education significantly reduced the magnitude of the HR, while the dietary factors had a negligible effect on the HR and its statistical significance. Tumors located at the border between the 241 242 cardia and the distal stomach (overlapping) or those whose localization was unknown had an HR of 243 1.43 (1.18, 1.73). These tumors could be either CGC or NCC and therefore the specific HR for this

category is not easily interpretable. Regarding histology, the HRs for the intestinal and diffuse types
were, respectively, 1.18 (1.03, 1.34) and 1.33 (1.06, 1.67), with no significant heterogeneity (*p*-value
0.31 for the Wald test).

The LR tests assessing departure from linearity were no statistically significant (for all GC as well 247 248 as for CGC and NCC) and therefore we assumed that the effects of ISD on risk can be reasonably well 249 represented by means of a linear dose-response relationship. Using this feature the increase (or decrease) in risk for any given value of the ISD it can be estimated (Supplemental Figure 2). For 250 251 instance, taking the median of the ISD as the reference (representing medium inflammatory potential of the diet in our population), the subjects with an ISD corresponding to the 10<sup>th</sup> percentile (assumed 252 to have a high anti-inflammatory diet) had a significant decrease in risk of GC of 27% (HR 0.73, CI 253 0.62, 0.85), and those with ISD corresponding to the 90<sup>th</sup> percentile (assumed to have a high pro-254 inflammatory diet) had a significant increased risk of 29% (HR 1.29, CI 1.13, 1.46). The 255 256 corresponding HR (95% CI) for CGC were 0.69 (0.51, 0.92) and 1.35 (1.07, 1.70), and 0.91 (0.70,

257 1.19) and 1.07 (0.87, 1.33) for NCC.

No significant differences in the association of GC, CGC or NCC with the ISD were observed 258 between men and women, according to age or by educational level (Table 3). Since tobacco smoking, 259 260 BMI and physical activity may contribute to low-grade chronic inflammation, we also explored whether the effect of the inflammatory potential of the diet on the risk of GC was modified by 261 262 smoking status and different levels of BMI and physical activity (Table 3). Although the association seemed to be more marked for smokers and subjects with normal weight, mainly for CGC, no 263 264 significant interactions were observed between the ISD and smoking status, BMI or physical activity level, either in all GCs or in tumors from the cardia or non-cardia regions. 265

The association between the ISD (for one SD increase of the ISD) and GC by country was assessed by means of a meta-analytic approach (**Figure 1**). All countries but Italy had HRs above the unity, although statistically significant estimates were observed only for UK, Sweden and Denmark (the countries with the largest number of cases). However these effects can be considered homogenous since the test of heterogeneity between countries according to a random effects model was not statistically significant. No heterogeneity was evident for CGC or NCC, although the patterns were less consistent, mainly for CGCs, owing to the small number of cases (detailed results in

# 273 Supplemental Table 3).

Finally, in order to assess the potential effect of reverse causality produced by a modification of the diet induced by a pre-existing (not clinically evident) condition, we excluded the subjects with follow-up below 2 years, which excluded 77 GC cases. In these analyses the adjusted HR (95%-CI) for each SD increase in the score was 1.22 (1.09, 1.37) for all GC, as compared with the 1.25 (1.12, 1.39) in the whole data set (**Table 2**). The corresponding estimates for the CGC and NCC were 1.29 (1.04, 1.60) and 1.05 (0.87, 1.27).

280

## 281 DISCUSSION

282 We have observed that the inflammatory potential of the diet, as measured by the ISD is associated 283 with higher risk of GC in a population of European adults. Each increase of one SD of the score significantly increased the GC risk by 25%; subjects eating a diet with the highest ISD (4<sup>th</sup> quartile) 284 have a 66% increase in GC risk as compared with those in the lowest quartile. This pattern seems to 285 be more consistent for tumors located in the cardia than for those located in the distal stomach, while 286 287 no differences were seen between the two major histological types (intestinal and diffuse). As far as we know this is the first prospective study on the association between GC risk and 288 289 inflammatory potential of the diet. Our results are consistent with those reported in an Italian hospital-290 based study with 230 GC cases and 547 matched controls (17). The adjusted odds ratio (OR) 291 comparing the highest to the lowest quartile of the DII was 2.35 (95% CI 1.32, 4.20), while in our population the adjusted HR was 1.66 (1.26, 2.20). This study did not provide results according to 292 anatomical tumor site of GC or by histological types. Although the results from both studies cannot be 293 294 directly compared as they are based upon different indexes, the DII and the ISD are actually close to 295 each other (see Methods). In our population the Pearson's correlation coefficient between the ISD and the DII was 0.91, with *p*-value <0.001. A population-based case-control study addressed the 296 relationship of DII and the risk of esophageal cancer in Sweden (16) reported separate results for 255 297 298 adenocarcinoma of the gastro-esophageal junction (GEJ) compared with 806 controls. The adjusted

OR (95% CI) comparing the fourth versus the first quartile of the DII was 2.04 (1.24, 3.36), similar to our estimate for the CGC (1.94, 95% CI 1.14, 3.30). The definition of GEJ has led to controversies and they have been alternatively considered as esophageal or gastric tumors, but in many instances they are still classified within the CGCs (26).

303 A role of inflammation in the pathogenesis of GCs has a strong biological plausibility (4-7), and 304 several dietary components have potential to modulate chronic inflammation (27). In previous studies we have shown a potential role in the GC risk of foods and nutrients that are in turn determinants of 305 306 the inflammatory potential of the diet measured by the ISD. For instance an increased risk of GC was 307 found to be associated with higher intake of red and processed meat, especially for NCC (28), while lower risks were associated with higher consumption of fruit and vegetables, mainly for CGC (29), 308 cereal fiber (30) and dietary flavonoids (31). A reduced GC risk was also observed with higher 309 plasmatic levels of vitamin C (32) as well as with higher circulating levels of some carotenoids, 310 311 retinol, and  $\alpha$ -tocopherol (33).

The inflammatory potential of the diet seems to have an independent effect on GC risk, not 312 313 explained by other factors. The multivariate model included a list of potential confounders selected a priori, based upon the known risk factors of CGC and NCC, and found to be associated with the 314 315 inflammatory potential of the diet in our population. Particular consideration was given to dietary factors as potential confounders; some of them (alcohol consumption, energy intake) are components 316 317 of the ISD, while other (intake of red and processed meat, intake citrus and non-citrus fresh fruit) are major sources of dietary included in the calculation of the ISD. On one side, including simultaneously 318 the ISD and the above mentioned factors in a model could produce overadjustment or collinearity; on 319 320 the other hand, these dietary factors may be true causes of GC and excluding them from the model 321 could result in effect estimates affected by residual confounding. We try to avoid these unwanted 322 effects by introducing into the multivariate model the residuals of a linear regression on the ISD of 323 each dietary variable (intakes of alcohol, energy, red and processed meat, and citrus and non-citrus fresh fruit) Therefore the HR of the ISD accounts for all the inflammatory potential of the diet, 324 325 whereas the HRs for each dietary factor account for their potential effect on GC by mechanisms other than inflammation. 326

327 Among the strengths of our study are the prospective design and its high statistical power, owing 328 to a large number of cases, an accurate case-ascertainment, and the ability to carry out specific 329 analyses according to histology and tumor localization of GC. The latter is particularly relevant since 330 there is growing evidence that CGC and NCC have different pathological and epidemiological 331 features. One of the most prominent is the differential role of *H. pylori*: chronic infection with *H*. 332 pylori is acknowledged as a cause of NCC (34), but no clear association with CGC. Moreover, recent studies have shown that eventually all cases of NCC have been previously infected by H. pylori, 333 334 suggesting that it is a necessary cause of this cancer (35). Therefore, it is unlikely that the association 335 between the ISD and NCC risk has been confounded by lack of adjustment by *H. pylori* infection. However, since the inflammatory process associated to *H. pylori* infection may be related with some 336 features of the bacterium such as virulence factors (36), detailed information on H. pylori infection 337 would have been useful to assess its potential modifying effect of the inflammatory potential of diet. 338 339 Although we have no data to assess this hypothesis, it could be that the inflammatory pathway leading to cancer in the distal part of the stomach if mostly driven by changes in the gastric mucosa induced 340 by *H. pylori* infection, and other factors related to chronic inflammation (i.e. diet) do not add too 341 much to already established process, while in the cardia, dietary factors (and maybe obesity) are more 342 343 relevant regarding the chronic inflammation associated with carcinogenesis.

344 A limitation of our study is that the estimation of the inflammatory potential of diet is based upon 345 the self-reported information on usual diet, gathered by means of methods relying on the subject's memory. Although we used validated tools (37) the potential for error measurement can never be 346 347 ruled out. In order to minimize the potential for measurement error in the usual diet subjects with implausible diets (those in the highest and lowest 1% of the distribution of the ratio between energy 348 349 intake and estimated energy requirement) were excluded; in addition a linear regression calibration 350 approach using data from 24hDR data was applied (25) and calibrated dietary intake was used to 351 calculate the ISD. On the other hand, since dietary information was collected on healthy individuals at 352 the beginning of the study, measurement errors would be expected to be non-differential. It is likely that some measurement error may persist; however, its effect would most likely dilute the true 353 354 association. Finally, we lack information on the usual consumption of anti-inflammatory drugs or

supplements, nor was information collected on foods preserved by salting or sodium intake; all thesefactors could have affected both the inflammatory potential and GC risk.

In summary, our results suggest that a diet with higher inflammatory potential is associated with 357 increased risk of GC; such association seems to be more consistent for gastric carcinomas located in 358 359 the cardia than for those located in the distal stomach. This effect seems to be independent of other risk factors of GC and other conditions related to chronic inflammation such as smoking, adiposity or 360 low levels of physical activity. They also suggest that beyond the potential effects of specific dietary 361 components, diet may play a role in gastric carcinogenesis as an overall modulator of low-grade 362 chronic inflammation. Further research including biomarkers of inflammation together with the 363 inflammatory potential of the diet would help to better understand the mechanisms underlying the role 364 365 of diet-related inflammation and gastric carcinogenesis.

366

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369	Authors' Contribution: AA designed and conducted the research, contributed to the data analysis,
370	wrote the manuscript and had primary responsibility for the final content of the manuscript. PJ
371	designed and conducted the research, contributed to the data analysis, and had primary responsibility
372	for the final content of the manuscript. VC and CB performed the statistical analysis. ER is the overall
373	coordinator of the EPIC study. All authors contributed to recruitment, data collection and acquisition,
374	biological sample collection, and follow-up and/or management of the EPIC cohort and to the
375	interpretation of the present findings and approval of the final version of the manuscript for
376	publication.

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			Gastric	(CGC /	ISD	
	Ν	(%)	cancer	NCC)	Median (p25, p75)	Mean (95% CI) <sup>1</sup>
Sex						
Men	142241	(29.9%)	509	(157 / 166)	-0.46 (-1.56, 0.61)	-0.30 (-0.31, -0.29)
Women	333919	(70.1%)	404	(79 / 175)	0.94 (-0.24, 1.93)	0.68 (0.67, 0.68)
Age at recruitment (years)						
< 40	56146	(11.8%)	25	(1 /13)	-0.43 (-1.98, 1.18)	-0.48 (-0.49, -0.47)
40 to <50	145768	(30.6%)	124	(32 / 50)	0.59 (-0.71, 1.73)	0.41 (0.41, 0.42)
50 to <60	181378	(38.1%)	388	(104 / 140)	0.59 (-0.56, 1.64)	0.51 (0.51, 0.52)
$\geq 60$	92868	(19.5%)	376	(99 / 138)	0.70 (-0.46, 1.75)	0.60 (0.59, 0.61)
Educational level						
None/primary not completed	20926	(4.4%)	55	(5 / 29)	0.43 (-1.04, 1.76)	0.14 (0.12, 0.16)
Primary	121856	(25.6%)	390	(77 / 166)	1.25 (0.09, 2.25)	1.10 (1.09, 1.10)
Technical/professional	105864	(22.2%)	218	(75 / 69)	0.60 (-0.68, 1.78)	0.48 (0.47, 0.49)
Secondary	97204	(20.4%)	94	(27 / 23)	0.60 (-0.49, 1.53)	0.34 (0.34, 0.35)
Longer (inc. university)	113379	(23.8%)	121	(38 / 47)	-0.32 (-1.51, 0.82)	-0.29 (-0.30, -0.28)
Unknown	16931	(3.6%)	35	(14 / 7)	0.01 (-1.30, 1.11)	-0.32 (-0.34, -0.30)
Smoking status						
Never	233096	(49.0%)	337	(60 / 151)	0.43 (-0.83, 1.48)	0.12 (0.12, 0.13)
Former	126822	(26.6%)	264	(81 / 81)	0.20 (-1.07, 1.43)	0.22 (0.21, 0.23)
Current	106564	(22.4%)	299	(91 / 105)	1.06 (-0.20, 2.18)	1.09 (1.08, 1.10)
Pipe/cigar/occasional/other <sup>2</sup>	9678	(2.0%)	13	(4 / 4)	1.40 (0.28, 2.24)	1.02 (0.99, 1.05)
BMI $(kg/m^2)$						
< 25.0	246060	(51.7%)	343	(81 / 120)	0.44 (-0.84, 1.53)	0.18 (0.18, 0.19)
25.0-29.9	166134	(34.9%)	397	(122 / 149)	0.53 (-0.73, 1.72)	0.55 (0.54, 0.56)
$\geq$ 30.0	63966	(13.4%)	173	(33 / 72)	0.87 (-0.44, 1.93)	0.71 (0.70, 0.72)
Alcohol consumption						. ,

 Table 1
 Main characteristics, number of events, and Inflammatory Score of the Diet (ISD) in the EPIC population.

non consumer	60724	(12.8%)	128	(17 / 51)	1.35 (0.27, 2.22)	0.85 (0.83, 0.86)
< 45.0 g/day	390277	(82.0%)	697	(194 / 261)	0.43 (-0.84, 1.57)	0.30 (0.29, 0.30)
45.0 - 59.9 g/day	12905	(2.7%)	39	(14 / 14)	-0.20 (-1.23, 0.87)	0.44 (0.41, 0.46)
$\geq 60.0$ g/day	12254	(2.6%)	49	(11 / 15)	-0.13 (-1.23, 0.95)	0.74 (0.72, 0.77)
Red meat (g/day, quartiles)						
< 16.11	119108	(25.0%)	167	(35 / 69)	0.22 (-1.48, 1.58)	-0.22 (-0.23, -0.21)
16.11-34.86	118974	(25.0%)	223	(49 / 96)	0.79 (-0.45, 1.85)	0.47 (0.46, 0.48)
34.87-63.10	119038	(25.0%)	244	(60 / 95)	0.73 (-0.43, 1.78)	0.62 (0.61, 0.63)
≥ 63.11	119040	(25.0%)	279	(92 / 81)	0.32 (-0.77, 1.34)	0.66 (0.65, 0.67)
Processed meat (g/day, quartiles)						
< 10.51	119040	(25.0%)	155	(35 / 51)	0.15 (-1.47, 1.48)	-0.24 (-0.25, -0.23)
10.51-24.25	119040	(25.0%)	211	(63 / 76)	0.72 (-0.47, 1.80)	0.40 (0.39, 0.41)
24.26-43.85	119063	(25.0%)	240	(53 / 93)	0.73 (-0.40, 1.77)	0.62 (0.61, 0.63)
≥ 43.86	119017	(25.0%)	307	(85 / 121)	0.39 (-0.75, 1.52)	0.76 (0.75, 0.77)
Citrus fruit (g/day, quartiles)						
< 8.23	121096	(25.4%)	268	(90 / 71)	0.99 (-0.22, 2.05)	0.81 (0.80, 0.82)
8.23-31.32	117206	(24.6%)	219	(61 / 83)	0.63 (-0.56, 1.71)	0.55 (0.54, 0.56)
31.33-70.52	119116	(25.0%)	185	(46 / 68)	0.46 (-0.75, 1.55)	0.28 (0.27, 0.29)
$\geq$ 70.53	118742	(24.9%)	241	(39 / 119)	-0.06 (-1.36, 1.22)	-0.11 (-0.12, -0.10)
Other fresh fruit (g/day, quartiles)						
< 64.46	119126	(25.0%)	231	(75 / 72)	1.08 (-0.03, 2.09)	0.98 (0.97, 0.99)
64.46-133.05	118954	(25.0%)	245	(68 / 87)	0.70 (-0.56, 1.77)	0.54 (0.54, 0.55)
133.06-226.10	119040	(25.0%)	221	(56 / 91)	0.36 (-0.87, 1.48)	0.19 (0.18, 0.20)
≥ 226.11	119040	(25.0%)	216	(37 / 91)	-0.13 (-1.40, 1.11)	-0.18 (-0.19, -0.17)

<sup>1</sup> Age, sex, and energy-adjusted means (95% CI) obtained from a linear regression model; the *p*-values comparing these means are always <0.001; for categorized variables with a

categories are based upon quantitative values (age, BMI, alcohol consumption, and all other dietary variables) this value corresponds to the *p*-trend.

<sup>2</sup> Includes occasional smokers, exclusive smokers of cigar and/or pipe, and smokers with unknown status and/or unknown amount smoked.

Abbreviations: CGC: cardia gastric cancer; NCC: non-cardia cancer.

**Table 2** Adjusted hazard ratios (HRs) and 95% CI of gastric cancer (by tumor subsite and histologic type) according to the Inflammatory Score of the Diet (ISD) in

 the EPIC population.

		ISD continuous <sup>1</sup>				
	Quartile 1	Quartile 2	Quartile 3	Quartile 4	<i>p</i> -trend	HR (95% CI)
Gastric cancer						
Basic model <sup>2</sup>	Referent	1.36 (1.11, 1.67)	1.77 (1.42, 2.20)	2.17 (1.70, 2.77)	< 0.001	1.38 (1.26, 1.52)
Multivariate model <sup>3</sup>	Referent	1.25 (1.02, 1.55)	1.50 (1.19, 1.89)	1.66 (1.26, 2.20)	< 0.001	1.25 (1.12, 1.39)
Cardia gastric cancer						
Basic model <sup>2</sup>	Referent	1.47 (0.99, 2.18)	2.03 (1.34, 3.07)	2.82 (1.79, 4.43)	< 0.001	1.51 (1.28, 1.80)
Multivariate model <sup>4</sup>	Referent	1.32 (0.88, 1.98)	1.64 (1.06, 2.55)	1.94 (1.14, 3.30)	0.011	1.30 (1.06, 1.59)
Non-cardia cancer						
Basic model <sup>2</sup>	Referent	1.23 (0.88, 1.72)	1.66 (1.16, 2.36)	1.52 (1.01, 2.28)	0.02	1.21 (1.04, 1.42)
Multivariate model <sup>5</sup>	Referent	1.17 (0.83, 1.65)	1.39 (0.96, 2.02)	1.07 (0.70, 1.70)	0.55	1.07 (0.89, 1.28)
Overlapping/unknown site						
Basic model <sup>2</sup>	Referent	1.41 (1.00, 2.00)	1.61 (1.10, 2.35)	2.44 (1.62, 3.68)	< 0.001	1.45 (1.24, 1.70)
Multivariate model <sup>3</sup>	Referent	1.31 (0.92, 1.87)	1.46 (0.98, 2.17)	2.35 (1.46, 3.77)	0.001	1.43 (1.18, 1.73)
GC, intestinal type						
Basic model <sup>2</sup>	Referent	1.36 (1.06, 1.75)	1.81 (1.40, 2.35)	2.25 (1.69, 3.01)	< 0.001	1.35 (1.21, 1.51)
Multivariate model <sup>3</sup>	Referent	1.26 (0.98, 1.62)	1.53 (1.16, 2.01)	1.65 (1.18, 2.31)	0.002	1.18 (1.03, 1.34)
GC, diffuse type						
Basic model <sup>2</sup>	Referent	1.33 (0.89, 1.99)	1.59 (1.02, 2.47)	1.71 (1.04, 2.82)	0.029	1.41 (1.16, 170)
Multivariate model <sup>3</sup>	Referent	1.18 (0.78, 1.78)	1.28 (0.81, 2.04)	1.30 (0.73, 2.31)	0.34	1.33 (1.06, 1.67)

<sup>1</sup>Hazard ratio (HR) per each increase in one standard deviation (SD) of the ISD.

<sup>2</sup>Stratified by age and center, and adjusted for sex and energy intake.

<sup>3</sup>Multivariate model: basic model and further adjusted by: educational level, tobacco smoking, BMI, alcohol consumption, and intake of red meat, processed meat, citrus fruit, and other fresh fruit (all the dietary variables expressed as residuals with respect to ISD).

<sup>4</sup>Multivariate model for cardia gastric cancers: model (3) excluding intake of red meat and processed meat.

<sup>5</sup>Multivariate model for non-cardia cancers: model (3) excluding BMI.

**Table 3** Association between the Inflammatory Score of the Diet (ISD) and gastric cancer risk by age, sex,

 education and non-dietary variables associated with chronic inflammation.

			HR (95% CI) <sup>1</sup>	
	-	Gastric cancer (GC) <sup>2</sup>	Cardia (CGC) <sup>3</sup>	Non-cardia (NCC) <sup>4</sup>
Sex	Men	1.22 (1.05, 1.41)	1.21 (0.95, 1.56)	1.12 (0.87, 1.46)
	Women	1.31 (1.09, 1.57)	1.53 (1.04, 2.25)	1.05 (0.80, 1.40)
<i>p</i> -value for in	teraction	0.54	0.16	0.8
Age at	<50 years	1.18 (0.88, 1.59)	1.37 (0.78, 2.39)	1.13 (0.74, 1.73)
recruitmnent	50 to <60 years	1.26 (1.06, 1.48)	1.46 (1.08, 1.98)	1.08 (0.82, 1.43)
	$\geq 60$ years	1.28 (1.08, 1.51)	1.17 (0.86, 1.60)	1.05 (0.79, 1.41)
<i>p</i> -value for in	teraction	0.57	0.43	0.28
Educational	None / Primary	1.22 (1.03, 1.45)	1.34 (0.91, 1.96)	0.91 (0.70, 1.18)
level	Technical/professional	1.22 (0.98, 1.52)	1.10 (0.77, 1.58)	1.33 (0.89, 2.00)
	Secondary	1.39 (0.95, 2.04)	1.52 (0.75, 3.08)	1.28 (0.62, 2.63)
	Longer (inc. university)	1.18 (0.89, 1.57)	1.32 (0.83, 2.09)	1.15 (0.69, 1.91)
<i>p</i> -value for in	teraction	0.45	0.84	0.15
Smoking	Never	1.02 (0.84, 1.25)	1.06 (0.70, 1.62)	0.99 (0.74, 1.33)
status	Former	1.40 (1.15, 1.71)	1.47 (1.04, 2.09)	0.97 (0.68, 1.38)
	Current	1.35 (1.11, 1.64)	1.41 (1.01, 1.97)	1.24 (0.89, 1.74)
<i>p</i> -value for in	teraction	0.82	0.68	0.81
BMI	< 25.0	1.47 (1.23, 1.77)	1.55 (1.10, 2.18)	1.13 (0.83, 1.55)
$(kg/m^2)$	25.0-29.9	1.15 (0.97, 1.36)	1.26 (0.94, 1.68)	1.03 (0.78, 1.37)
	$\geq 30.0$	1.07 (0.82, 1.39)	0.95 (0.54, 1.67)	1.09 (0.70, 1.69)
<i>p</i> -value for interaction		0.14	0.32	0.28
Physical	Inactive	1.38 (1.09, 1.73)	1.11 (0.71, 1.73)	1.22 (0.83, 1.78)
activity	Moderately inactive	1.25 (1.00, 1.56)	1.53 (0.99, 2.34)	1.14 (0.77, 1.68)
	Moderately active	1.29 (1.01, 1.66)	2.15 (1.36, 3.41)	0.80 (0.54, 1.21)
	Active	1.12 (0.89, 1.42)	1.08 (0.73, 1.61)	1.06 (0.72, 1.57)
<i>p</i> -value for interaction		0.41	0.54	0.71

<sup>1</sup>Hazard ratio (HR) per each increase in one standard deviation (SD) of the ISD.

<sup>2</sup>Stratified by age and center, and adjusted for sex and energy intake, educational level, tobacco smoking, BMI, alcohol

consumption, and intake of red meat, processed meat, citrus fruit, and other fresh fruit (all the dietary variables

expressed as residuals with respect to the ISD).

<sup>3</sup>As model for GC excluding red and processed meat intake.

<sup>4</sup>As model for GC excluding BMI.

The p-value for interaction is based upon the Likelihood ratio (LR) test

## Legends for figures

Figure 1 Association between the Inflammatory Score of the Diet (ISD) and gastric cancer in EPIC by country.

### Footnote:

HR (95% CI): Hazard ratio for each increase of one standard deviation of the ISD, estimated from a Cox model stratified by age and center, and adjusted for sex, energy intake, educational level, tobacco smoking, BMI, alcohol consumption, and intake of red meat, processed meat, citrus fruit, and other fresh fruit (all the dietary variables as residuals with respect to the ISD).

RE Model: summary estimate from a random effects meta-analysis

Heterogeneity test:  $Q_{(9 df)} = 7.35$ , *p*-value 0.60